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BUSINESS CYCLE IMPLICATIONS OF INTERNAL CONSUMPTION HABIT FOR NEW KEYNESIAN MODELS

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Abstract

We study the implications of internal consumption habit for new Keynesian dynamic stochastic general equilibrium (NKDSGE) models. Bayesian Monte Carlo methods are employed to evaluate NKDSGE model fit. Simulation experiments show that internal consumption habit often improves the ability of NKDSGE models to match the spectra of output and consumption growth. Nonetheless, the fit of NKDSGE models with internal consumption habit is susceptible to the sources of nominal rigidity, to spectra identified by permanent productivity shocks, to the choice of monetary policy rule, and to the frequencies used for evaluation. These vulnerabilities indicate that the specification of NKDSGE models is fragile.

Key Words: Consumption Habit; New Keynesian; Propagation; Monetary Transmission; Posterior Predictive Analysis; Bayesian Monte Carlo.

JEL Classification Number: E10, E20, E32.

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1. INTRODUCTION

This paper evaluates the impact of internal consumption habit on the empirical fit of new Keynesian dynamic stochastic general equilibrium (NKDSGE) models. Consumption habit is often superinduced in DSGE models to improve fit, but the role consumption habit has in a NKDSGE model is not settled. For example, Del Negro, Schorfheide, Smets, and Wouters (2007) find that consumption habit contributes to a NKDSGE model matching the hump-shaped response of output to a Taylor rule shock, but Christiano, Eichenbaum, and Evans (2005) do not; instead they use a money growth shock.

This is not the case for real business cycle (RBC) models. Consumption habit is known to be successful at closing the distance between RBC models and aggregate quantity and asset price moments since the work of Boldrin, Christiano, and Fisher (2001).¹ An explanation for this success, suggested by Eichenbaum and Hansen (1990) and Heaton (1995), is that consumption habit is a real friction that imposes costs when a household substitutes utility intertemporally.

We present an experiment using a log linearized Euler equation that studies the effects the utility costs of internal consumption habit have in inducing intertemporal complementarity in consumption growth from date t to date $t+j$, $j > 1$, instead of intertemporal substitution in the level of consumption between dates t and $t+1$. The experiment shows that as internal consumption habit increases in utility, the household reacts to a positive real interest rate shock by pushing the peak response in its consumption growth from date $t+1$ to date $t+j$. The experiment also documents that frequency by frequency the variance of consumption growth falls steadily across the entire spectrum, holding internal consumption habit fixed. As internal consumption habit increases, the spectral density (SD) of consumption growth exhibits less and less amplitude frequency by frequency.²

Beyond this experiment, the implications of internal consumption habit for NKDSGE models are studied using several population moments. Our analysis relies on population moments because, as Geweke (2010) argues, confronting the predictive density of a DSGE model with a sufficiently large vector of sample moments almost surely negates it. Geweke (2010) proposes the minimal econometric interpretation (MEI) of DSGE models to address this problem. Under the MEI, priors on the DSGE model yield a distribution for population moments, but neither for the actual data (*i.e.*, the model's likelihood)

1. Consumption habit is first grafted into a growth model by Ryder and Heal (1973). Nason (1988), Sundareshan (1989), and Constantinides (1990) are early attempts at solving risk-free rate and equity premium puzzles with consumption habit. Pollak (1976) shows that long-run utility with linear habit describes long-run behavior rather than long-run preferences. Rozen (2010) gives an axiomatic treatment of linear intrinsic habit. An excellent survey of consumption habit in macro and finance is Schmitt-Grohé and Uribe (2007); also see Nason (1997).

2. Otrok, Ravikumar, and Whiteman (2002) make a similar point using spectral utility analysis.

nor for sample moments. We employ the MEI to evaluate 12 NKDSGE models on prior population moments and posterior population moments generated by estimating Bayesian vector autoregressions (BVARs), adapting Bayesian Monte Carlo methods developed by DeJong, Ingram, and Whiteman (1996), Geweke (1999), and McCausland (2004).³

This paper evaluates the fit of NKDSGE models with and without internal consumption habit on permanent and transitory *SDs* of output and consumption growth. Our choice of these moments is guided by earlier studies of the U.S. business cycle. Among others, Galí (1991) finds that the *SD* of consumption growth is not flat, which violates the permanent income hypothesis (PIH). Cogley and Nason (1995b) observe that RBC models cannot reproduce the *SD* of output growth because it peaks in the business cycle frequencies. They show, along with Nason and Cogley (1994), that DSGE models fail to duplicate output's response to permanent and transitory shocks. Thus, we explore NKDSGE model fit on moments known to have power to judge competing theories of the U.S. business cycle.

The identification of permanent and transitory *SDs* relies on output and consumption being orthogonal to transitory nominal shocks in the long run. This long-run monetary neutrality (LRMN) restriction is used to identify posterior population *SDs* given estimates of unrestricted BVARs and prior population versions of the same moments that are generated by habit and non-habit NKDSGE models. A NKDSGE model is evaluated by comparing its prior moments to the posterior moments. The comparison relies on VARs identified by LRMN because these empirical models are used to compute prior moments on a NKDSGE model's synthetic samples and posterior moments on the sample data.

We also study the impact various combinations of nominal frictions, monetary policy rules, and internal consumption habit have on NKDSGE model fit. Motivation is provided by Christiano, Eichenbaum, and Evans (CEE), Smets and Wouters (2007), and Dupor, Han, and Tsai (2009). CEE report that a money growth shock is transmitted by sticky wages, but not by sticky prices. In contrast, Smets and Wouters present evidence that sticky prices and wages have about the same effect on NKDSGE model fit given that monetary policy is a Taylor rule. However, Dupor, Han, and Tsai obtain estimates

3. The MEI eschews estimation, but we recognize that estimating NKDSGE models with limited information estimators (LIEs) and classical and Bayesian likelihood methods is useful. Nonetheless, estimation does not guarantee problem-free evaluation of NKDSGE models. For example, model misspecification is an issue for frequentists when evaluating NKDSGE models. Applying a LIE to an incorrectly specified NKDSGE model yields estimates that are not robust to changing identification (or instruments), as suggested by Dupor, Han, and Tsai (2009). A strength of Bayesian likelihood-based estimation is that it updates the priors of NKDSGE model parameters, which is absent from the MEI of Geweke. Not updating priors can produce misleading measures of model fit when these priors are badly constructed, but posterior distributions created by Bayesian likelihood-based estimation suffer similar problems from poorly formed priors as discussed by Poirier (1998). Given these issues, we view the MEI as providing analysis that complements studies that estimate NKDSGE models.

of a NKDSGE model identified by a productivity shock, instead of a Taylor rule shock, that point to flexible prices and durability in consumption rather than habit.

The Bayesian Monte Carlo experiments show that internal consumption habit improves the fit of NKDSGE models to the posterior *SDs* of output and consumption growth. However, the fit is sensitive to (1) changes in the mix of nominal rigidities, (2) the choice of monetary policy rule, (3) matching *SDs* on permanent total factor productivity (TFP) shocks, and (4) using the entire spectrum rather than just the business cycle frequencies. These vulnerabilities reveal that the specification of NKDSGE models is fragile.

The rest of the paper is constructed as follows. Section 2 discusses internal consumption habit and NKDSGE models. Our approach to NKDSGE model evaluation is outlined in section 3. Results appear in section 4. Section 5 concludes.

2. INTERNAL CONSUMPTION HABIT AND NKDSGE MODELS

This section describes household preferences with internal consumption habit, studies the internal consumption habit propagation mechanism, and sketches the baseline NKDSGE model.

2.1 *Internal consumption habit*

Internal habit operates on lagged household consumption, unlike external habit, which assumes that lags of aggregate consumption appear in utility; see Abel (1990).⁴ Household preferences are intertemporally separable and separable across (net) consumption, labor disutility, and real balances,

$$\mathcal{U}\left(c_t, c_{t-1}, n_t, \frac{H_t}{P_t}\right) = \ln[c_t - hc_{t-1}] - \frac{\gamma}{1+\gamma} n_t^{1+\frac{1}{\gamma}} + \ln\left[\frac{H_t}{P_t}\right], \quad (1)$$

where c_t , n_t , h , γ , H_t , and P_t are household consumption, household labor supply, the internal consumption habit parameter $h \in (0, 1)$, the Frisch labor supply elasticity ($\gamma > 0$), household cash at the end of date $t-1$, and the aggregate price level, respectively. Since internal consumption habit ties the choice of c_t to c_{t-1} , the marginal utility of consumption is forward-looking,

$$\lambda_t = \frac{1}{c_t - hc_{t-1}} - \mathbf{E}_t\left\{\frac{\beta h}{c_{t+1} - hc_t}\right\},$$

assuming $0 < c_t - hc_{t-1}$ for all t , where the household discount factor $\beta \in (0, 1)$ and $\mathbf{E}_t\{\cdot\}$ is the

4. The appendix shows there is an observational equivalence for multiplicative internal and external consumption habit using the onto mapping from additive to multiplicative consumption habit parameters that Dennis (2009) constructs. This indicates little generality is lost by focusing on internal consumption habit. The appendix is available at <http://hermes-ir.lib.hit-u.ac.jp/rs/bitstream/10086/23297/1/070econDP12-08.pdf>.

mathematical expectation operator given date t information.

2.2 The internal consumption habit propagation mechanism

Forward-looking marginal utility suggests internal habit acts as a propagation mechanism for consumption. Intuition about this propagation mechanism is grounded in the Euler equation

$$\lambda_t = \beta \mathbf{E}_t \left\{ \frac{\lambda_{t+1} R_{t+1}}{1 + \zeta_{t+1}} \right\},$$

where R_t is the nominal rate and $1 + \zeta_{t+1}$ ($= P_{t+1}/P_t$) is date $t+1$ inflation.⁵ A log linear approximation of the Euler equation and the marginal utility of consumption give a second-order stochastic difference equation for demeaned consumption growth, $\widetilde{\Delta c}_t$, whose solution is

$$\widetilde{\Delta c}_t = \varphi_1 \widetilde{\Delta c}_{t-1} + \frac{(\alpha^* - \beta h)(\alpha^* - h)}{\alpha^2 \varphi_2} \sum_{j=0}^{\infty} \varphi_2^{-j} \mathbf{E}_t \widetilde{q}_{t+j}, \quad (2)$$

where the stable and unstable roots are $\varphi_1 = h\alpha^{*-1}$ and $\varphi_2 = \alpha^*(\beta h)^{-1}$, α^* is steady-state growth, the demeaned real interest rate is $\widetilde{q}_t = \widetilde{R}_t - \frac{\zeta^*}{1 + \zeta^*} \widetilde{\zeta}_t$, and ζ^* is mean inflation.⁶

Equation (2) yields the response of $\widetilde{\Delta c}_t$ to a shock in \widetilde{q}_t by calibrating it, β , and α^* . The calibration to quarterly data sets \widetilde{q}_t to a first-order autoregression, AR(1), with AR1 coefficient $\rho_q = 0.87$ and $[\beta \alpha^*]' = [0.993 \exp(0.004)]'$.⁷ Impulse response functions (IRFs) and SDs of $\widetilde{\Delta c}_t$ are produced using equation (2), the calibration, $h = [0.0 \ 0.15 \ 0.35 \ 0.50 \ 0.65 \ 0.85]$, and a positive unit shock in \widetilde{q}_t .

The top window of figure 1 plots IRFs across $h = [0.0 \ 0.15 \ 0.35 \ 0.50 \ 0.65 \ 0.85]$. As h increases from 0.15 to 0.85, IRFs become more humped shaped with the peak in $\widetilde{\Delta c}_t$ moving from a horizon of one quarter (the plain blue solid line) to quarter 6 (the purple solid line with the \times symbol). Subsequently, $\widetilde{\Delta c}_t$ steadily falls, but the decay rate slows as h rises. These IRFs contrast with the IRF generated when household preferences are not infected by internal consumption habit. The top window of figure 1 depicts the latter IRF with a (black) solid line with the symbol \diamond . This IRF is not hump shaped and, because it is equivalent to the autocorrelation function of \widetilde{q}_t , its slow decay is governed by powers of ρ_q .⁸ Hence, internal habit by switching consumption from an intertemporal substitute at $h = 0$ to a complement, given $h \in (0, 1)$, creates an economically meaningful propagation mechanism.

5. Consumption-based asset pricing models with habit and local substitution through service are estimated by Eichenbaum and Hansen (1990) and Heaton (1995). They reject the adjustment cost hypothesis in favor of services flows. However, the data support habit if local substitutability operates at lower frequencies than the sampling frequency of consumption.

6. The appendix constructs equation (2), assuming a unit root TFP shock drives trend consumption. Also, \widetilde{R}_t is the demeaned nominal federal funds rate, while ζ_t is set to the growth rate of the implicit GDP deflator.

7. The SIC selects an AR(1) for \widetilde{q}_t over any lag length up to 10 on a 1954Q1–2002Q4 sample. See the appendix for details.

8. The equivalence stems from a linear approximate Euler equation under $h = 0$ equating $\widetilde{\Delta c}_t$ to \widetilde{q}_t up to an Euler error.

This internal consumption habit propagation mechanism is similar to the one discussed by CEE using the NKDSGE models they estimate. Their baseline NKDSGE model, in which h is estimated to be about 0.65, generates a hump-shaped response in output that peaks six quarters after an innovation to a nominal shock. The top panel of figure 1 includes an IRF created using $h = 0.65$, the (aquamarine) dot-dashed plot, that has similar dynamics with a peak at quarter 5. This shows that equation (2) can produce an IRF that resembles those found using estimated NKDSGE models.

Another way to understand the role of internal consumption habit in consumption growth dynamics is the SD of $\widetilde{\Delta c}_t$. The SD of $\widetilde{\Delta c}_t$, which we denote as $SD_{\widetilde{\Delta c}}$, decomposes the variance of $\widetilde{\Delta c}_t$ frequency by frequency. The bottom panel of figure 1 shows that the $SD_{\widetilde{\Delta c}}$ s have the greatest power at the long run (*i.e.*, frequency zero), which indicates that the lowest frequencies contribute most to the variance of $\widetilde{\Delta c}_t$. The $SD_{\widetilde{\Delta c}}$ indexed by no habit exhibits more power from frequency zero to about four years per cycle compared to the $SD_{\widetilde{\Delta c}}$ s driven by internal consumption habit. Over these frequencies, power declines monotonically for the $SD_{\widetilde{\Delta c}}$ as h steps up from 0.15 to 0.85. Thus, greater internal consumption habit lowers the power of $SD_{\widetilde{\Delta c}}$ from the lowest into the business cycle frequencies.

To summarize, internal consumption habit makes it more costly for the household to smooth utility intertemporally. This dictates intertemporal complementarity, which persuades the household to push the peak in its consumption growth response to a shock further into the future. These utility costs also produce SD s with less power in the lower frequencies as h rises. The rest of this paper uses this intuition to study the implications of internal consumption habit for NKDSGE models.

2.3 A new Keynesian DSGE model

Besides internal consumption habit, the baseline NKDSGE model contains (a) capital adjustment costs, (b) variable capital utilization, (c) fully indexed Calvo-staggered wage setting by monopolistic households with heterogeneous labor supply, and (d) fully indexed Calvo-staggered price setting by monopolistic final goods firms. Elements (a)-(c) contribute to the optimization problems faced by households with addresses, $\ell \in [0, 1]$, on the unit circle. The budget constraint of household ℓ is

$$\frac{H_{t+1}}{P_t} + \frac{B_{t+1}}{P_t} + c_t + x_t + a(u_t)k_t + \tau_t = r_t u_t k_t + \frac{W_t(\ell)}{P_t} n_t(\ell) + \frac{H_t}{P_t} + R_t \frac{B_t}{P_t} + \frac{D_t}{P_t}, \quad (3)$$

where B_{t+1} is the stock of government bonds carried from date t into date $t + 1$, x_t is investment, k_t is household capital at the end of date $t - 1$, τ_t is a lump-sum government transfer, r_t is the real rental rate of k_t , $W_t(\ell)$ is the nominal wage of household ℓ , R_t is the nominal return on B_t , D_t is dividends received

from firms, $u_t \in (0, 1)$ is the capital utilization rate, and $a(u_t)$ is its cost function. A change in u_t forces household ℓ to forgo $a(\cdot)$ units of consumption per unit of capital. The investment adjustment costs specification, which is adapted from CEE, is placed in the law of motion of household capital

$$k_{t+1} = (1 - \delta)k_t + \left[1 - S\left(\frac{1}{\alpha} \frac{x_t}{x_{t-1}}\right)\right] x_t, \quad \delta \in (0, 1), \quad 0 < \alpha, \quad (4)$$

where δ is the capital depreciation rate and $\alpha (= \ln \alpha^*)$ is deterministic TFP growth. The cost function $S(\cdot)$ is strictly convex, where $S(1) = S'(1) = 0$ and $S''(1) \equiv \varpi > 0$.

Given k_0 , B_0 , and c_{-1} , the expected discounted lifetime utility function of household ℓ

$$\mathbf{E}_t \left\{ \sum_{i=0}^{\infty} \beta^i \mathbf{U} \left(c_{t+i}, c_{t+i-1}, n_{t+i}(\ell), \frac{H_t}{P_t} \right) \right\} \quad (5)$$

is maximized by choosing c_t , k_{t+1} , H_{t+1} , B_{t+1} , and $W_t(\ell)$ subject to period utility (1), budget constraint (3), the law of motion of capital (4), and downward-sloping labor demand.

Households charge firms $W_t(\ell)$ per unit of differentiated labor services in a monopolistic market in which a Calvo-staggered nominal wage mechanism operates. Given θ is the wage elasticity, the labor supply aggregator is $N_t = \left[\int_0^1 n_t(\ell)^{(\theta-1)/\theta} d\ell \right]^{\theta/(\theta-1)}$. Labor market monopoly imposes downward-sloping labor demand schedules, $n_t(\ell) = \left[W_t / W_t(\ell) \right]^\theta N_t$, on firms, where the nominal wage index is $W_t = \left[\int_0^1 W_t(\ell)^{1-\theta} d\ell \right]^{1/(1-\theta)}$ and its aggregator is $W_t = \left[(1 - \mu_W) W_{c,t}^{1-\theta} + \mu_W (\alpha^* \zeta_{t-1} W_{t-1})^{1-\theta} \right]^{1/(1-\theta)}$. This Calvo-staggered nominal wage mechanism has households updating their nominal wage to $W_{c,t}$ at probability $1 - \mu_W$. At probability μ_W , households receive the date $t-1$ nominal wage indexed by steady-state TFP growth, α^* , and ζ_{t-1} . In this case, the optimal nominal wage condition is

$$\left[\frac{W_{c,t}}{P_{t-1}} \right]^{1+\theta/\gamma} = \left(\frac{\theta}{\theta-1} \right) \frac{\mathbf{E}_t \sum_{i=0}^{\infty} \left[\beta \mu_W \alpha^{*-\theta(1+1/\gamma)} \right]^i \left[\left[\frac{W_{t+i}}{P_{t+i-1}} \right]^\theta N_{t+i} \right]^{1+1/\gamma}}{\mathbf{E}_t \sum_{i=0}^{\infty} \left[\beta \mu_W \alpha^{*(1-\theta)} \right]^i \lambda_{t+i} \left[\frac{W_{t+i}}{P_{t+i-1}} \right]^\theta \left[\frac{P_{t+i}}{P_{t+i-1}} \right]^{-1} N_{t+i}}. \quad (6)$$

Equation (6) smooths nominal wage growth, which forces labor supply to absorb TFP and monetary policy shocks. Since shifts in labor supply alter production and intra- and intertemporal margins, shock volatility and persistence are realized, for example, as output and consumption fluctuations.

Monopolistically competitive firms produce final goods that households consume. The consumption aggregator is $c_t = \left[\int_0^1 \mathcal{Y}_{D,t}(j)^{(\xi-1)/\xi} dj \right]^{\xi/(\xi-1)}$, where $\mathcal{Y}_{D,t}(j)$ is household final good demand for the output of a firm with address j on the unit interval. The j th final good firm aims to meet this

demand with its output, $y_t(j)$, by mixing capital, $K_t(j)$, rented and labor, $N_t(j)$, hired from households net of fixed cost N_0 given labor-augmenting TFP, A_t , in the constant returns to scale technology, $[u_t K_t(j)]^\psi [N_t(j) - N_0] A_t^{1-\psi}$, $\psi \in (0, 1)$. Fixed labor cost N_0 satisfies the needs of monopolistic competition in the final goods market. For the NKDSGE model to have a permanent shock, TFP is a random walk with drift, $A_t = A_{t-1} \exp\{\alpha + \varepsilon_t\}$, with its innovation, $\varepsilon_t \sim \mathcal{N}(0, \sigma_\varepsilon^2)$.

Firm j maximizes profits by choosing its price $P_t(j)$, subject to $y_{D,t}(j) = [P_t/P_t(j)]^\xi Y_{D,t}$, where ξ is the price elasticity, $Y_{D,t}$ is aggregate demand, and the price index is $P_t = \left[\int_0^1 P_t(j)^{1-\xi} \right]^{1/(1-\xi)}$. Under Calvo-staggered price setting, firms update their optimal price $P_{c,t}$ at probability $1-\mu_p$. Or with probability μ_p , firms are stuck with P_{t-1} scaled by ζ_{t-1} , which defines the price aggregator $P_t = \left[(1-\mu_p)P_{c,t}^{1-\xi} + \mu_p (\zeta_{t-1}P_{t-1})^{1-\xi} \right]^{1/(1-\xi)}$. The firm's problem yields the optimal forward-looking price

$$\frac{P_{c,t}}{P_{t-1}} = \left(\frac{\xi}{\xi-1} \right) \frac{\mathbf{E}_t \sum_{i=0}^{\infty} (\beta\mu_p)^i \lambda_{t+i} \phi_{t+i} Y_{D,t+i} \zeta_{t+i}^\xi}{\mathbf{E}_t \sum_{i=0}^{\infty} (\beta\mu_p)^i \lambda_{t+i} Y_{D,t+i} \zeta_{t+i}^{\xi-1}} \quad (7)$$

of a firm able to update its price. Inflation is smoothed by equation (7). The same responses are induced in output and consumption in response to TFP and monetary policy shocks because of the reaction of monopolistically competitive firms to variation in the aggregate price level.

We close the NKDSGE model with one of two monetary policy rules. CEE identify monetary policy with a $MA(\infty)$ money growth process. It is equivalent to the AR(1) money growth rule (MGR)

$$m_{t+1} = (1 - \rho_m)m^* + \rho_m m_t + \mu_t, \quad |\rho_m| < 1, \quad \mu_t \sim \mathcal{N}(0, \sigma_\mu^2), \quad (8)$$

where $m_{t+1} = \ln(M_{t+1}/M_t)$ and m^* is its mean. A model using the MGR (8) is labeled NKDSGE-MGR. The mnemonic NKDSGE-TR refers to models closed by the Taylor rule (TR)

$$(1 - \rho_R \mathbf{L})R_t = (1 - \rho_R) \left(R^* + a_\zeta \mathbf{E}_t \zeta_{t+1} + a_{\tilde{Y}} \tilde{Y}_t \right) + v_t, \quad |\rho_R| < 1, \quad v_t \sim \mathcal{N}(0, \sigma_v^2), \quad (9)$$

where $R^* = \exp(m^* - \alpha)/\beta$. The TR (9) assumes the Taylor principle, $1 < a_\zeta$ is obeyed, $0 < a_{\tilde{Y}}$, and expected inflation, $\mathbf{E}_t \zeta_{t+1}$, and transitory output, \tilde{Y}_t , are computed without measurement errors.

The government finances B_t , interest on B_t , and a lump-sum transfer τ_t with new bond issuance $B_{t+1} - B_t$, lump-sum taxes τ_t , and money creation, $M_{t+1} - M_t$. These sources and uses of funds give the government the budget constraint $P_t \tau_t = [M_{t+1} - M_t] + [B_{t+1} - (1 + R_t)B_t]$. We assume government

debt is in zero net supply, $B_{t+1} = 0$ and $P_t \tau_t = M_{t+1} - M_t$, along the equilibrium path at all dates t .

The decentralized economy requires goods, labor, and money markets to clear in equilibrium. Equilibrium has $K_t = k_t$ given $0 < r_t$, $N_t = n_t$ given $0 < W_t$, and $M_t = H_t$ given $0 < P_t, R_t$. The aggregate resource constraint, $Y_t = C_t + I_t + a(u_t)K_t$, follows, where $C_t = c_t$ and $I_t = x_t$.⁹

3. BAYESIAN MONTE CARLO STRATEGY

We employ posterior and prior population moments to judge the fit of 12 variants of the NKDSGE model. Bayesian Monte Carlo simulations are run that combine sample data, a BVAR, its priors, a Markov chain Monte Carlo (MCMC) simulator, LRMN, and the Blanchard and Quah (1989) decomposition to create structural vector moving averages (SVMAs) from which posterior distributions of *SDs* of output growth, ΔY , and consumption growth, ΔC , are constructed. Prior distributions of *SDs* of ΔY and ΔC are approximated using SVMAs generated by applying LRMN and the Blanchard and Quah (BQ) decomposition to VARs estimated by least squares on synthetic data that are simulated from NKDSGE models with parameters drawn from independent priors. The SVMAs serve as the econometric link between posterior and prior *SDs* and the sample data. The multidimensional posterior and prior *SDs* are collapsed into scalar Kolmogorov-Smirnov (*KS*) goodness of fit statistics to compute the confidence interval criterion (*CIC*) of DeJong, Ingram, and Whiteman (1996). A *CIC* measures the overlap of posterior and prior *KS* statistic distributions. We engage the *CIC* to quantify NKDSGE model fit.

3.1 Output and consumption moments

This section describes the process we use to construct the permanent and transitory *SDs* of ΔY and ΔC . These *SDs* are grounded on SVMAs just-identified by a LRMN restriction embedded in the NKDSGE model and the orthogonality of permanent and transitory shock innovations. Given these restrictions, SVMAs are computed by imposing the BQ decomposition on second-order VARs, VAR(2)s, and BVAR(2)s. When a VAR(2) is estimated on synthetic NKDSGE model samples of $[\Delta \ln Y_t \ \Delta \ln P_t]'$ or $[\Delta \ln C_t \ \Delta \ln P_t]'$, the SVMAs produce prior permanent and transitory *SDs* of ΔY or ΔC . The posterior versions of these *SDs* are created from SVMAs based on BVAR(2)s derived from sample data and our priors. The BQ decomposition identifies the SVMAs because the TFP innovation ε_t is the permanent shock and the transitory shock is either the MGR innovation μ_t or TR innovation ν_t .¹⁰ The identified

9. A rational expectations equilibrium equates, on average, firm and household subjective forecasts of r_t , and A_t to the objective outcomes produced by the decentralized economy. The list includes μ_t and R_t under the MGR (8) or ν_t for the TR (9). A flexible price regime (spot labor market) adds P_t (W_t).

10. The appendix of Blanchard and Quah (1989) includes a theorem that establishes necessary and sufficient conditions under

$\Delta \ln Y_t - \Delta \ln P_t$ ($\Delta \ln C_t - \Delta \ln P_t$) system recovers a SVMA that has a vertical long-run aggregate supply (PIH) curve and a serially correlated “output (consumption) gap.”

As an example consider the SVMA

$$\begin{bmatrix} \Delta \ln Y_t \\ \Delta \ln P_t \end{bmatrix} = \sum_{j=0}^{\infty} \mathbb{B}_j \begin{bmatrix} \varepsilon_{t-j} \\ u_{t-j} \end{bmatrix}, \quad \text{where } \mathbb{B}_j = \begin{bmatrix} \mathbb{B}_{\Delta Y, \varepsilon, j} & \mathbb{B}_{\Delta Y, u, j} \\ \mathbb{B}_{\Delta P, \varepsilon, j} & \mathbb{B}_{\Delta P, u, j} \end{bmatrix}, \quad (10)$$

that equates the monetary policy shock with the innovation u_t of the Taylor rule (9). Elements of \mathbb{B}_j are just-identified by imposing (i) orthogonality on ε_t and u_t and (ii) the LRMN restriction $\mathbb{B}_{\Delta Y, u}(\mathbf{1}) = 0$ (i.e., $\ln Y_{t+j}$ is independent of u_t at the infinite horizon, $j \rightarrow \infty$). These assumptions decompose the SVMA (10) into univariate SMA(∞)s, $\mathbb{B}_{\Delta Y, \varepsilon}(\mathbf{L})\varepsilon_t$ and $\mathbb{B}_{\Delta Y, u}(\mathbf{L})u_t$, for ΔY . The former (latter) SMA(∞) is the IRF of ΔY with respect to the permanent shock ε_t (transitory shock u_t).

We grab the SMA(∞) of $\mathbb{B}_{\Delta Y, \varepsilon}(\mathbf{L})\varepsilon_t$ and $\mathbb{B}_{\Delta Y, u}(\mathbf{L})u_t$ from the SVMA (10) to compute permanent and transitory *SDs* of ΔY . Since the SVMA (10) is also a Wold representation of $[\Delta \ln Y_t \ \Delta \ln P_t]'$, its *SD* (at frequency ω) is $SD_{[\Delta Y \ \Delta P]}(\omega) = (2\pi)^{-1} \Gamma_{[\Delta Y \ \Delta P]} \exp(-i\omega)$, where $\Gamma_{[\Delta Y \ \Delta P]}(l) = \sum_{j=0}^{\infty} \mathbb{B}_j \mathbb{B}'_{j-l}$. The convolution $\Gamma_{[\Delta Y \ \Delta P]}(l)$ is expanded at horizon j to obtain

$$\mathbb{B}_j \mathbb{B}'_{j-l} = \begin{bmatrix} \mathbb{B}_{\Delta Y, \varepsilon, j} \mathbb{B}_{\Delta Y, \varepsilon, j-l} + \mathbb{B}_{\Delta Y, u, j} \mathbb{B}_{\Delta Y, u, j-l} & \mathbb{B}_{\Delta Y, \varepsilon, j} \mathbb{B}_{\Delta P, \varepsilon, j-l} + \mathbb{B}_{\Delta Y, u, j} \mathbb{B}_{\Delta P, u, j-l} \\ \mathbb{B}_{\Delta P, \varepsilon, j} \mathbb{B}_{\Delta Y, \varepsilon, j-l} + \mathbb{B}_{\Delta P, u, j} \mathbb{B}_{\Delta Y, u, j-l} & \mathbb{B}_{\Delta P, \varepsilon, j} \mathbb{B}_{\Delta P, \varepsilon, j-l} + \mathbb{B}_{\Delta P, u, j} \mathbb{B}_{\Delta P, u, j-l} \end{bmatrix}.$$

The matrix's off-diagonal entries are elements of the cross-covariance function of ΔY and ΔP and, therefore, map to co- and quad-spectra. The autocovariance function of ΔY with respect to ε_t (u_t) is the upper-left diagonal, $\mathbb{B}_{\Delta Y, \varepsilon, j} \mathbb{B}_{\Delta Y, \varepsilon, j-l}$ ($\mathbb{B}_{\Delta Y, u, j} \mathbb{B}_{\Delta Y, u, j-l}$). We exploit the SMAs $\mathbb{B}_{\Delta Y, \varepsilon}(\mathbf{L})\varepsilon_t$ and $\mathbb{B}_{\Delta Y, u}(\mathbf{L})u_t$ that are along the diagonal, to parameterize permanent and transitory *SDs* of ΔY at frequency ω

$$SD_{\Delta Y, l}(\omega) = \frac{1}{2\pi} \left| \mathbb{B}_{\Delta Y, l, 0} + \mathbb{B}_{\Delta Y, l, 1} e^{-i\omega} + \mathbb{B}_{\Delta Y, l, 2} e^{-i2\omega} + \dots + \mathbb{B}_{\Delta Y, l, j} e^{-ij\omega} + \dots \right|^2, \quad l = \varepsilon, u,$$

given the BQ decomposition assumption that the structural shock innovations have unit variances.

Before computing $SD_{\Delta Y, l}(\omega)$, we truncate its polynomial at $j = 40$, or a 10-year horizon.¹¹

3.2 The moments to match: Posterior distributions

We engage MCMC software of Geweke (1999) and McCausland (2004) to create posterior distri-

which bivariate ARs identify the correct responses to a permanent shock and a transitory shock when the truth is that there are several of these shocks. The BQ decomposition is satisfied, according to the theorem, when responses, say, of ΔY and ΔP to permanent or transitory shocks are equivalent up to a scalar lag operator. Since the shocks found in NKDSGE models are often AR(1)s, the theorem predicts that adding these shocks to a NKDSGE model does not create spurious identification.

11. This approach to estimating *SDs* extends the ideas of Akaike (1969) and Parzen (1974).

butions of SVMAs. These posterior distributions consist of $J = 5000$ SVMA parameter vectors that are grounded on BVAR(2)s, our priors, a 1954Q1–2002Q4 sample ($T = 196$) of $\Delta \ln Y_t$, $\Delta \ln C_t$, and $\Delta \ln P_t$, LRMN, and the BQ decomposition.¹² These J vectors are used to calculate distributions of posterior permanent and transitory SD s of ΔY and ΔC . We label these posterior moments $SD_{P,\Delta Y}$ and $SD_{P,\Delta C}$.

3.3 Bayesian simulation methods II: Prior distributions

It takes multiple steps to solve and simulate the NKDSGE models. Since these models have a permanent TFP shock, stochastic detrending of optimality and equilibrium conditions is needed before log-linearizing around deterministic steady states. We engage an algorithm of Sims (2002) to solve for log linear approximate equilibrium laws. Synthetic samples are created by feeding sequences of the shock innovations ε_t and μ_t or ν_t into the equilibrium laws of motion given initial conditions and draws from the priors of NKDSGE model parameters. The appendix describes these procedures in detail.

Table 1 lists the priors of the NKDSGE models. These priors reflect our uncertainty about NKDSGE model parameters. For example, the first item listed in table 1 is the consumption habit parameter h . We give h an uninformative prior that is drawn from a uniform distribution with end points 0.05 and 0.95 in table 1. The uninformative prior reflects a belief that any $h \in [0.05, 0.95]$ is as likely as another. Non-habit NKDSGE models are defined by the degenerate prior $h = 0$.

Priors are also taken from earlier DSGE model studies.¹³ We set the means of the priors of $[\beta \ \delta \ \alpha \ \psi]'$ = $[0.9930 \ 0.0200 \ 0.0040 \ 0.3500]'$ that are consistent with the Cogley and Nason (1995b) calibration. Uncertainty about $[\beta \ \gamma \ \delta \ \alpha \ \psi]'$ is captured by 95 percent coverage intervals, which contain values in Nason and Cogley (1994) and Hall (1996). We equate the prior of the investment cost of adjustment parameter ϖ to estimates reported by Bouakez, Cardia, and Ruge–Murcia (2005).¹⁴ The standard deviation of TFP shock innovations, σ_ε , is endowed with a uniform prior because the DSGE literature suggests that any draw from $[0.0070, 0.0140]$ is equally likely. The source of the prior mean of the Frisch labor supply elasticity, $\gamma = 1.55$, and its 95 percent coverage interval of $[0.5, 2.6]$ is Kimmel

12. The Geweke-McCausland MCMC software treats the BVARs as Bayesian linear multivariate (*i.e.*, seemingly unrelated) regressions. The prior of the intercept and slope parameters of the BVARs are independent of the prior on the inverse of the covariance (*i.e.*, precision) matrix of the regression errors in the Geweke-McCausland MCMC code. This restriction yields a multivariate normal (Wishart) marginal distribution for the prior of the intercept and slope parameters (covariance matrix). An empirical Bayes prior (*i.e.*, maximum likelihood estimates) is our choice for providing the priors on the intercept, slope, and covariance matrix parameters with numerical content. We give the Wishart distribution prior on the covariance matrix five degrees of freedom. The Geweke-McCausland MCMC code also employs a Gibbs sampler to draw from the posterior that generates updates of the latent variables using a first-order Markov process; see section 2.1 of McCausland (2004) for more details.

13. Several priors are centered on sample means of the consumption-output ratio, labor input, federal funds rate, and inflation using the 1954Q1–2002Q4 sample. We also fix $N_0 = 0.1678$ and $r^* = 1.0050$.

14. The capacity utilization rate, u_t , and its cost function, $a(\cdot)$, are fixed at the steady state, $u^* = 1$ and $a(1) = 0$. Determinate solutions are achieved by setting $a''(1)/a'(1) = 1.174$.

and Kniesner (1998). They estimate labor supply functions on U.S. panel data.

There are four sticky price and wage parameters to calibrate. The prior means are $[\xi \ \mu_P \ \theta \ \mu_w]'$ = $[12.0 \ 0.55 \ 15.0 \ 0.7]'$. The mean of ξ implies a steady-state price markup, $\xi/(\xi - 1)$, of 9 percent with a 95 percent coverage interval that runs from 5 to 33 percent. This coverage interval blankets estimates found in Basu and Fernald (1997) and CEE. More uncertainty surrounds the priors of μ_P , θ , and μ_w . Sbordone (2002), Nason and Slotsve (2004), Lindé (2005), and CEE suggest a 95 percent coverage interval for μ_P of $[0.45, 0.65]$. Likewise, a 95 percent coverage interval of $[0.04, 0.25]$ suggests substantial uncertainty around the 7 percent prior mean household wage markup, $\theta/(\theta - 1)$. The degenerate mean of μ_w and its 95 percent coverage interval reveal stickier nominal wages than prices, as found by CEE and Rabanal and Rubio-Ramírez (2005), but we imbue it with greater uncertainty.

The money growth rule (8) is calibrated to estimates from a 1954Q1–2002Q4 sample of the monetary base. The point estimates are degenerate priors for $[m^* \ \rho_m \ \sigma_\mu]'$ = $[0.011 \ 0.628 \ 0.006]'$. We give these prior means less precision than found for the sample estimates. For example, the lower end of the 95 percent coverage interval of ρ_m is near 0.46. CEE note that $\rho_m \approx 0.5$ implies the AR(1) money growth rule (8) mimics the persistence of their MA(∞) money growth shock policy process.

The calibration of the interest rate rule (9) obeys the Taylor principle and $a_y \in (0, 1)$. The degenerate prior of a_π is 1.80. We assign a small role to movements in transitory output, \tilde{Y} , with a prior mean of 0.05 for a_y . The 95 percent coverage intervals of a_π and a_y rely on estimates reported by Smets and Wouters (2007). The interest rate rule (9) is also calibrated to smooth R_t given a prior mean of 0.65 and a 95 percent coverage interval of $[0.55, 0.74]$ that incorporates estimates found in Guerron-Quintana (2010). Ireland (2001) is the source of the prior mean of the standard deviation of the monetary policy shock, $\sigma_u = 0.0051$, and its 95 percent coverage interval, $[0.0031, 0.0072]$. We assume all shock innovations are uncorrelated at leads and lags (*i.e.*, $\mathbf{E}\{\varepsilon_{t+i} u_{t+q}\} = 0$, for all i, q).

Prior draws of NKDSGE model parameters are applied to its log linear approximation to generate a synthetic sample of length $\mathcal{M} = \mathcal{W} \times T$. We set $\mathcal{W} = 5$ to approximate prior population distributions.¹⁵

After estimating VAR(2)s on these samples, LRMN and the BQ decomposition are applied to construct

15. Geweke's MEI compares posterior population distributions to prior population distributions. For the latter distributions, this requires $\mathcal{M} \rightarrow \infty$, which implies $\mathcal{W} > 1$. When $\mathcal{W} = 1$, the model evaluation exercise falls under the rubric of prior predictive analysis. Since distributions of prior population SDs can only be approximated by simulating the NKDSGE models, the quality of the approximation is checked by simulating the NKDSGE models for $\mathcal{W} = 2, 3, \dots$. At $\mathcal{W} = 5$, the change in the mean prior SDs and the measures of NKDSGE model fit (described in section 3.4) is small (*i.e.*, about 0.5^{-6}). This gives us confidence that $\mathcal{M} = 5 \times 196$ is a sufficiently long sample to approximate $\mathcal{M} \rightarrow \infty$.

SVMAs.¹⁶ The SVMAs are employed to compute prior population permanent and transitory SDs of ΔY and ΔC . Since our uncertainty about the theory — the parameters of the NKDSGE models — is embedded in these prior population moments, the subscript \mathcal{T} is used to denote $SD_{\mathcal{T},\Delta Y}$ and $SD_{\mathcal{T},\Delta C}$.

3.4 Measures of fit

An intermediate step in measuring NKDSGE model fit is to construct a KS goodness of fit statistic similar to one used by Cogley and Nason (1995a). They evaluate the fit of a RBC model to the $SD_{\Delta Y}$ with the KS . Since multidimensional SDs are mapped into a scalar KS statistic, it is an efficient metric for summarizing the information in SDs when conducting a model evaluation exercise.

Bayesian Monte Carlo experiments produce distributions of permanent and transitory $SD_{\mathcal{P},\Delta Y}$, $SD_{\mathcal{P},\Delta C}$, $SD_{\mathcal{T},\Delta Y}$, and $SD_{\mathcal{T},\Delta C}$. We convert the multidimensional SDs into scalar posterior and prior KS statistics, $KS_{\mathcal{P}}$ and $KS_{\mathcal{T}}$. The first step normalizes permanent and transitory $SD_{\mathcal{P},\Delta Y}$, $SD_{\mathcal{P},\Delta C}$, $SD_{\mathcal{T},\Delta Y}$, and $SD_{\mathcal{T},\Delta C}$ on sample permanent and transitory SDs of ΔY or ΔC , $\widehat{SD}_{\Delta Y}$ and $\widehat{SD}_{\Delta C}$. The sample SDs are calculated using a SVMA estimated on the actual data with length T . Define the normalization $\mathcal{R}_{\mathcal{D},j}(\omega) = \widehat{SD}(\omega)/SD_{\mathcal{D},j}(\omega)$ at replication j and frequency ω , where $\mathcal{D} = \mathcal{P}, \mathcal{T}$. Next, compute the partial sum $\mathcal{V}_{\mathcal{D},j}(2\pi q/\mathcal{H}) = 2\pi\mathcal{H}^{-1} \sum_{\ell=1}^q \mathcal{R}_{\mathcal{D},j}(2\pi\ell/\mathcal{H})$, where $\mathcal{H} = T$ when $\mathcal{D} = \mathcal{P}$ and $\mathcal{H} = \mathcal{M}$ otherwise. The partial sums are used to construct the partial difference $\mathcal{B}_{\mathcal{D},j}(\kappa) = 0.5\pi^{-1}\sqrt{2\mathcal{H}} \left[\mathcal{V}_{\mathcal{D},j}(\kappa\pi) - \kappa\mathcal{V}_{\mathcal{D},j}(\pi) \right]$, where $\kappa \in [0, 1]$ indicates $\mathcal{B}_{\mathcal{D},j}(\kappa)$ is evaluated on the entire spectrum. The scalar $KS_{\mathcal{D}}$ statistic at replication j is calculated as the maximal absolute value of $\mathcal{B}_{\mathcal{D},j}(\kappa)$, $KS_{\mathcal{D},j} = \text{Max}_{\kappa \in [0,1]} \left| \mathcal{B}_{\mathcal{D},j}(\kappa) \right|$. Thus, a $KS_{\mathcal{P},j}$ ($KS_{\mathcal{T},j}$) statistic measures the distance between a $SD_{\mathcal{P}}$ ($SD_{\mathcal{T}}$) and \widehat{SD} .

We engage the CIC to measure the fraction of J elements of a $KS_{\mathcal{T}}$ distribution that occupies an interval defined by lower and upper quantiles of the associated $KS_{\mathcal{P}}$ distribution at a $1-p$ percent confidence level.¹⁷ Suppose a habit NKDSGE model generates a distribution of the permanent $SD_{\mathcal{T},\Delta Y}$ that fills this interval with more than 30 percent of its elements ($CIC > 0.3$), but the corresponding non-habit model does not ($CIC \leq 0.3$).¹⁸ In this case (as DeJong, Ingram, and Whiteman (1996) imply in their analysis of RBC models), the habit NKDSGE model is a more plausible match to the permanent $SD_{\mathcal{P},\Delta Y}$. Note that this match requires the $SD_{\mathcal{T}}$ distribution to be “near” the $SD_{\mathcal{P}}$ distribution at several

16. The appendix constructs the map from SVAR to SVMA. The SVARs satisfy the invertibility condition of Fernández-Villaverde, Rubio-Ramírez, Sargent, and Watson (2007) across all Bayesian Monte Carlo simulations of the 12 NKDSGE models.

17. As in DeJong, Ingram, and Whiteman (1996), the CIC of a $KS_{\mathcal{T}}$ statistic distribution is set to $\frac{1}{1-p} \int_a^b KS_{\mathcal{T},j} dj$ for a $1-p$ percent confidence level, where a (b) is the lower $0.5p$ (upper $1-0.5p$) quantile and is normalized by $1-p$ to equal $\int_a^b KS_{\mathcal{P},j} dj$.

18. If the CIC of a NKDSGE model with internal consumption habit is larger than a NKDSGE model without it, the implication is that the Bayes factor supports the former model.

frequencies for a $KS_{\mathcal{T}}$ statistic distribution to cover more than 30 percent of the distribution of a $KS_{\mathcal{P}}$ statistic. Thus, a CIC constitutes a “joint test” of NKDSGE model fit.

The $KS_{\mathcal{P}}$ and $KS_{\mathcal{T}}$ statistics are computed on the entire spectrum (*i.e.*, the long- to the short-run) and the business cycle horizons of eight to two years per cycle. By isolating the business cycle fluctuations, we build on an insight of Diebold, Ohanian, and Berkowitz (1998). They focus on the business cycle frequencies when model misspecification corrupts measurement of short- and long-run fluctuations. We address this issue by limiting κ to frequencies of between eight and two years per cycle, $\kappa \in [0.064, 0.25]$, when compiling $KS_{\mathcal{P}}$ and $KS_{\mathcal{T}}$ distributions. This mitigates discounting NKDSGE models that perform well at business cycle horizons, but poorly on lower growth and higher short-run frequencies.

4. HABIT AND NON-HABIT NKDSGE MODEL EVALUATION

Before evaluating the fit of the NKDSGE models to posterior permanent and transitory $SD_{\mathcal{P},\Delta Y}$ s and $SD_{\mathcal{P},\Delta C}$ s, we study the means of these posterior moments. Figure 2 plots mean posterior permanent and transitory $SD_{\mathcal{P},\Delta Y}$ and $SD_{\mathcal{P},\Delta C}$. The mean $SD_{\mathcal{P},\Delta Y}$ ($SD_{\mathcal{P},\Delta C}$) decomposes variation in the average response of ΔY (ΔC) to permanent and transitory shocks frequency by frequency.¹⁹ The top (bottom) panel of figure 2 contains mean permanent (transitory) $SD_{\mathcal{P},\Delta Y}$ and $SD_{\mathcal{P},\Delta C}$. Mean $SD_{\mathcal{P},\Delta Y}$ appear as solid (blue) lines in figure 2 and mean $SD_{\mathcal{P},\Delta C}$ are depicted with (blue) ‘◆’ symbols.

Mean permanent $SD_{\mathcal{P},\Delta Y}$ and $SD_{\mathcal{P},\Delta C}$ display the greatest power at frequency zero in the top panel of figure 2. Subsequently, these SD s decay from the growth (*i.e.*, more than eight years per cycle) into the highest frequencies (*i.e.*, less than two years per cycle), which indicate these frequencies matter less and less for variation in the permanent components of ΔY and ΔC . Compared to the $SD_{\mathcal{P},\Delta C}$, the mean permanent $SD_{\mathcal{P},\Delta Y}$ exhibits about five times the power at the long run. This is consistent with the PIH and other business cycle theories that predict greater volatility in ΔY than in ΔC . However, the PIH is not supported by the mean permanent $SD_{\mathcal{P},\Delta C}$ because it is not flat across the entire spectrum.

The lower panel of figure 2 presents mean transitory $SD_{\mathcal{P},\Delta Y}$ and $SD_{\mathcal{P},\Delta C}$ with disparate shapes. The latter SD peaks around six years per cycle. Rather than a peak, the mean transitory $SD_{\mathcal{P},\Delta C}$ plateaus from the growth frequencies to four years per cycle before dropping off in the high frequencies. Thus, the transitory or “output gap” component of ΔY exhibits greater periodicity in the business cycle frequencies compared to the transitory component of ΔC .

19. A mean $SD_{\mathcal{P}}$ is computed across an ensemble of $SD_{\mathcal{P},j}$, $j = 1, \dots, J$, pointwise or frequency by frequency.

Figure 2 summarizes the challenge for NKDSGE models. The models must produce permanent $SD_{\mathcal{T},\Delta Y}$ and $SD_{\mathcal{T},\Delta C}$ that are not white noise, but instead have most power at the lowest frequencies, to match the permanent $SD_{\mathcal{P}}$ s displayed in the top panel of figure 2. The lower panel of figure 2 challenges NKDSGE models to generate transitory $SD_{\mathcal{T},\Delta Y}$ ($SD_{\mathcal{T},\Delta C}$) with the greatest power at the business cycle (growth) frequencies. As a result, NKDSGE models need empirically and economically meaningful propagation and monetary transmission mechanisms to match permanent and transitory $SD_{\mathcal{P}}$.

4.1 Quantifying NKDSGE model fit: CIC

Table 2 presents the *CIC* that measures the overlap of distributions of $KS_{\mathcal{P}}$ and $KS_{\mathcal{T}}$ statistics. The *CIC* is calculated at a 95 percent confidence level, $p = 0.05$. The overlap of these distributions gauges the fit of 12 NKDSGE models to permanent and transitory $SD_{\mathcal{P},\Delta Y}$ and $SD_{\mathcal{P},\Delta C}$ distributions.

The 12 NKDSGE models are defined by different combinations of nominal frictions as well as by monetary policy rule. The first NKDSGE model is our baseline that includes sticky prices and wages. From this baseline, two more NKDSGE models are created by stripping out one or the other nominal rigidities.²⁰ Baseline, sticky price (SPrice), and sticky wage (SWage) NKDSGE models have household preferences with either no consumption habit, $h = 0$, or internal consumption habit, $h \in [0.05, 0.95]$. These six NKDSGE models are doubled by defining monetary policy with the MGR (8) or the TR (9).

The *CIC* of baseline, SPrice, and SWage non-habit NKDSGE models is listed in the top panel of table 2. The lower panel contains the *CIC* of habit NKDSGE models. The NKDSGE-MGR and NKDSGE-TR models define monetary policy with the MGR (8) and the TR (9), respectively. Columns titled $[0, \pi]$ and $[8, 2]$ report *CIC* quantifying the overlap of distributions of $KS_{\mathcal{P}}$ and $KS_{\mathcal{T}}$ statistics on the entire frequency domain (*i.e.*, the long-run or frequency zero to the short-run) and the business cycle frequencies that run from eight to two years per cycle, respectively.

Table 2 contains the *CIC* that indicates internal consumption habit improves the fit of NKDSGE models. Habit NKDSGE models generate 18 $CIC > 0.3$ (or 37.5 percent) out of a possible 48 in the bottom half of table 2, but the top half of the table shows that non-habit NKDSGE models are responsible for only nine $CIC > 0.3$ (or 18.75 percent). Thus, internal consumption habit contributes to permanent and transitory $SD_{\mathcal{T},\Delta Y}$ and $SD_{\mathcal{T},\Delta C}$ that better replicate the posterior moments.

A striking feature of table 2 is the impact of the monetary policy rules on the fit of NKDSGE

20. The sticky wage NKDSGE model requires the degenerate prior $\mu_p = 0$ with fixed markup $\phi = (\xi - 1)/\xi$. When the nominal wage is flexible, households set their optimal wage period by period in sticky price NKDSGE models. In this case, the markup in the labor market is fixed at $(\theta - 1)/\theta$, which equals $n^{-1/\gamma}$, given $\mu_w = 0$.

models to transitory $SD_{P,\Delta Y}$ and $SD_{P,\Delta C}$. In these models, the MGR (8) has few successes at transmitting its shock innovation μ_t into ΔY and ΔC fluctuations that resemble the posterior moments. According to the CIC of table 2, only the SPrice habit NKDSGE-MGR model accomplishes this task and only on the business cycle frequencies of eight to two years per cycle.

The TR (9) improves NKDSGE model fit compared with the MGR (8). Of the 27 $CIC > 0.3$ found in table 2, the TR (9) is tied to 22. In the top half of table 2, eight of the nine $CIC > 0.3$ are linked to non-habit NKDSGE-TR models. Similarly, habit NKDSGE-TR models yield 14 of the 18 $CIC > 0.3$ in the lower half of table 2. Table 2 also shows that, on the business cycle frequencies of eight to two years per cycle, habit and non-habit NKDSGE-TR models match transitory $SD_{P,\Delta Y}$ and $SD_{P,\Delta C}$ with $CIC > 0.3$ in all 12 possible cases. Thus, table 2 gives evidence that habit and non-habit NKDSGE models have empirically useful monetary transmission mechanisms when initiated by the TR shock innovation ν_t . However, only habit NKDSGE-TR models replicate the transitory SD_{P} s on the entire spectrum because these models are responsible for 11 of 12 possible $CIC > 0.33$ in the lower half of table 2.

The propagation mechanisms of the NKDSGE models are not held in similar regard by the permanent $SD_{P,\Delta Y}$ and $SD_{P,\Delta C}$. These SD_{P} s are matched only when permanent $SD_{T,\Delta Y}$ and $SD_{T,\Delta C}$ are generated by SPrice habit and non-habit NKDSGE-MGR and -TR models. The bottom half of table 2 shows that a $CIC > 0.3$ appears in four of the four possible cases when SPrice habit NKDSGE-MGR and -TR models are asked to propagate the TFP innovation shock ε_t into the business cycle frequencies of eight to two years per cycle. The SPrice non-habit NKDSGE-MGR and -TR models are less successful at this task. Only the permanent $SD_{P,\Delta Y}$ is duplicated by these models. When the propagation mechanisms of these models are examined using the entire spectrum, a $CIC > 0.3$ is produced only for the permanent $SD_{P,\Delta Y}$. Nonetheless, this shows that measuring model fit on the entire spectrum leads to otherwise empirically relevant NKDSGE models being undervalued.

In summary, internal consumption habit confers a superior fit on NKDSGE models according to the CIC of table 2. The improved fit of habit NKDSGE models is predicated, in part, on the TR (9), but non-habit NKDSGE-TR models possess empirically and economically credible monetary transmission mechanisms when asked to match transitory SD_{P} s only on the business cycle frequencies of eight to two years per cycle. Limiting the analysis to the business cycle frequencies improves the fit of SPrice habit NKDSGE-MGR and -TR models when the match is to permanent $SD_{P,\Delta Y}$ and $SD_{P,\Delta C}$.²¹

21. The appendix reports results of Bayesian Monte Carlo experiments that substitute the Cramer-von Mises goodness of fit

4.2 NKDSGE model dynamics: Internal consumption habit

The habit NKDSGE models provide a superior fit to the permanent and transitory $SD_{P,\Delta Y}$ s and $SD_{P,\Delta C}$ s, according to the *CIC* of table 2. Although moment matching is useful for assessing the fit of the NKDSGE models, this evaluation process is not informative about the propagation and monetary transmission mechanism dynamics of NKDSGE models. This section and the next explore NKDSGE model dynamics by comparing mean permanent and transitory SD_{P} s and SD_{T} s.

We summarize evidence about NKDSGE model dynamics in figures 3 to 5. The figures consist of four rows and two columns of panels containing means of SD_{P} s and SD_{T} s. From top to bottom, the rows of figures 3 to 5 plot mean permanent $SD_{\Delta Y}$ s, transitory $SD_{\Delta Y}$ s, permanent $SD_{\Delta C}$ s, and transitory $SD_{\Delta C}$ s. The NKDSGE-MGR models generate SD_{T} s that appear in the left column of figures 3 to 5. The right column of these figures includes SD_{T} s that are produced by NKDSGE-TR models. Figures 3 to 5 display mean SD_{P} s with solid (blue) lines, mean SD_{T} s created by non-habit NKDSGE models with (green) dotted lines, and mean SD_{T} s tied to habit NKDSGE models with (red) dot-dashed lines. The mean SD_{T} s of baseline, SPrice, and SWage NKDSGE models are depicted in figures 3, 4, and 5, respectively.

Figures 3 to 5 display mean permanent and transitory $SD_{T,\Delta Y}$ s and $SD_{T,\Delta C}$ s, with (green) dotted plots in figures 3 to 5, produced by non-habit NKDSGE models that often exhibit greater power than the (red) dot-dash mean SD_{T} s created by habit NKDSGE models. As a result, habit NKDSGE models induce SD_{T} s that are either closer to or intersect mean SD_{P} s, especially at the business cycle frequencies.

These mean SD_{T} s reproduce an important feature of the $SD_{\Delta C}^{\sim}$ plotted in the bottom window of figure 1. Recall that the $SD_{\Delta C}^{\sim}$ are computed using the linear approximate Euler equation (2) and from top to bottom are indexed by $h = 0$ to 0.85. The bottom window of figure 1 shows that the $SD_{\Delta C}^{\sim}$ s converge pointwise to the horizontal axis as h increases. Since the mean permanent and transitory $SD_{T,\Delta Y}$ s and $SD_{T,\Delta C}$ s of figures 3 to 5 often have the same ordering, the utility adjustment costs inherent in internal consumption habit are responsible, in part, for propagation and monetary transmission in NKDSGE models that replicate the mean permanent and transitory $SD_{P,\Delta Y}$ s and $SD_{P,\Delta C}$ s of figure 2.

4.3 NKDSGE model dynamics: Mixing nominal frictions and monetary policy rules

This section explores the impact that different combinations of sticky prices, sticky wages, and

statistic for the *KS* statistic, estimate VAR(4)s instead of VAR(2)s, and replace the uniform prior, $h \sim U(0.05, 0.95)$, with either $h \sim U(0.05, 0.499)$, $h \sim U(0.50, 0.95)$, or $h \sim \beta(0.65, 0.15)$. The latter prior implies a 95 percent coverage interval for h of [0.38, 0.88]. These experiments reinforce the message table 2 has for the impact of internal consumption habit on NKDSGE model fit as well as the vulnerabilities of the fit of NKDSGE models to different combinations of nominal frictions and monetary policy rules.

monetary policy rule have on propagation and monetary transmission in NKDSGE models. For example, the first and third rows of figures 3 to 5 establish that there is no mix of nominal rigidities and monetary policy rule in habit or non-habit NKDSGE models that yield a credible facsimile of the mean permanent $SD_{P,\Delta Y}$ and $SD_{P,\Delta C}$. The large gaps between these posterior moments and mean permanent $SD_{T,\Delta Y}$ s and $SD_{T,\Delta C}$ s are not because NKDSGE models lack powerful propagation mechanisms. Instead, the top row of windows in figures 3 to 5 show that the 12 NKDSGE models propagate the TFP shock innovation ε_t into $SD_{T,\Delta Y}$ that often have greatest variation in the business cycle frequencies between eight and four years per cycle. In comparison, the mean permanent $SD_{P,\Delta Y}$ has greatest power at the long run before dropping off in the business cycle frequencies. Similarly, the third row of figures 3 to 5 displays gaps between the $SD_{P,\Delta C}$ and $SD_{T,\Delta C}$ s, which are often largest in the business cycle frequencies.

There is no NKDSGE model with inflation smoothing generated by the optimal forward-looking price setting of wage equation (7) that provides an economically meaningful monetary transmission mechanism. Figures 3 and 4 are clear that habit and non-habit NKDSGE models with sticky prices have difficulties reproducing mean transitory $SD_{P,\Delta Y}$ and $SD_{P,\Delta C}$. For example, the baseline and SPrice habit and non-habit NKDSGE-MGR models often generate power in mean transitory $SD_{T,\Delta Y}$ s and $SD_{T,\Delta C}$ s at the business cycle frequencies from eight to two years a cycle, which is excessive compared to the mean transitory $SD_{P,\Delta Y}$ s and $SD_{P,\Delta C}$ s. The distance between SD_{P} s and SD_{T} s is repaired only in part by switching from the MGR (8) to the TR (9) in the baseline and SPrice NKDSGE models. Only the bottom-right window of figure 3 reports that the baseline habit NKDSGE-TR model generates a mean transitory $SD_{T,\Delta C}$ with less power in the business frequencies, which moves it closer to the $SD_{P,\Delta C}$ s.

The mean transitory $SD_{T,\Delta Y}$ and $SD_{T,\Delta C}$ of figure 5 show that the nominal wage smoothing generated by the optimal forward-looking equation (6) contributes, along with internal consumption habit and the TR (9), to economically credible monetary transmission. This is the SWage habit NKDSGE model, whose mean transitory $SD_{T,\Delta Y}$ and $SD_{T,\Delta C}$ are depicted with (red) dot-dash plots in the even numbered rows of figure 5's right-hand column. Observe in the second row of figure 5 that the mean $SD_{T,\Delta Y}$ peaks between eight and four years per cycle, which echoes the shape of the mean transitory $SD_{P,\Delta Y}$. The SWage habit NKDSGE-TR model also yields a mean transitory $SD_{T,\Delta C}$ in the bottom-right window of figure 5 that reproduces the plateau in the growth frequencies of the mean transitory $SD_{P,\Delta C}$. The left-hand column of figure 5 shows that dropping internal consumption habit or swapping the MGR (8) for the TR (9) pushes mean transitory SD_{T} s away from the mean posterior moments.

5. Conclusion

This paper studies the business cycle implications of internal consumption habit for NKDSGE models. We examine the fit of 12 NKDSGE models that have different combinations of internal consumption habit, Calvo-staggered prices, and nominal wages, along with several other real rigidities. The NKDSGE models are confronted with posterior *SDs* of output and consumption growth identified by permanent TFP and transitory monetary policy shocks.

The fit of NKDSGE models with and without internal consumption habit is explored by comparing posterior population moments to theoretical prior population moments. Our analysis shows that the fit of NKDSGE models with consumption habit is susceptible to (1) changing the mix of nominal rigidities, (2) identifying *SDs* on permanent TFP shocks instead of transitory monetary policy shocks, and (3) evaluating *SDs* on the entire spectrum rather than the business cycle frequencies.

These results indicate vulnerabilities in the specification of NKDSGE models. Not unexpectedly, the new Keynesian monetary transmission mechanism is not the issue. There are combinations of sticky prices, sticky wages, and monetary policy rule that match the posterior *SDs* of output and consumption growth. Nonetheless, only when internal consumption habit, sticky nominal wages, and a Taylor rule are included in a NKDSGE model does it transmit monetary policy shocks into empirically plausible mean *SDs* of output and consumption growth. The vulnerabilities in NKDSGE model fit are tied to the mix of nominal rigidities and judging fit on the entire spectrum when the moment-matching exercise is identified with the permanent TFP shock. Thus, the economic and empirical relevance of the propagation mechanisms of NKDSGE models remain open to more research. We hope this paper plays a part in inspiring further research into the role real and nominal rigidities play in propagation as well as monetary transmission in NKDSGE models.

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TABLE 1: BAYESIAN CALIBRATION OF NKDSGE MODELS

		Prior Distribution	Mean	Standard Deviation	95 Percent Cover Interval
h	Internal Consumption Habit	Uniform	—	—	[0.0500, 0.9500]
β	H'hold Subjective Discount	Beta	0.9930	0.0020	[0.9886, 0.9964]
γ	Labor Supply Elasticity	Normal	1.5500	0.5360	[0.4995, 2.6005]
δ	Depreciation Rate	Beta	0.0200	0.0045	[0.0122, 0.0297]
α	Deterministic Growth Rate	Normal	0.0040	0.0015	[0.0011, 0.0064]
ϖ	Capital Adjustment Costs	Normal	4.7710	1.0260	[2.7601, 6.7819]
ψ	Capital's Share of Output	Beta	0.3500	0.0500	[0.2554, 0.4509]
σ_ε	TFP Growth Shock Std.	Uniform	—	—	[0.0070, 0.0140]
ξ	Final Good Dmd Elasticity	Normal	12.0000	4.0820	[3.9994, 20.0006]
μ_P	No Price Change Probability	Beta	0.5500	0.0500	[0.4513, 0.6468]
θ	Labor Demand Elasticity	Normal	15.0000	3.0800	[8.9633, 21.0367]
μ_W	No Wage Change Probability	Beta	0.7000	0.0500	[0.5978, 0.7931]
m^*	$\Delta \ln M$ Mean	Beta	0.0114	0.0030	[0.0063, 0.0180]
ρ_m	$\Delta \ln M$ AR1 Coef.	Beta	0.6278	0.0800	[0.4653, 0.7767]
σ_μ	$\Delta \ln M$ Shock Std.	Beta	0.0064	0.0012	[0.0043, 0.0090]
a_ζ	Taylor Rule $E_t \zeta_{t+1}$ Coef.	Normal	1.8250	0.2300	[1.3742, 2.2758]
$a_{\hat{Y}}$	Taylor Rule \hat{Y}_t Coef.	Normal	0.1000	0.0243	[0.0524, 0.1476]
ρ_R	Taylor Rule AR1 Coef.	Beta	0.6490	0.0579	[0.5317, 0.7578]
σ_v	Taylor Rule Shock Std.	Beta	0.0051	0.0016	[0.0025, 0.0087]

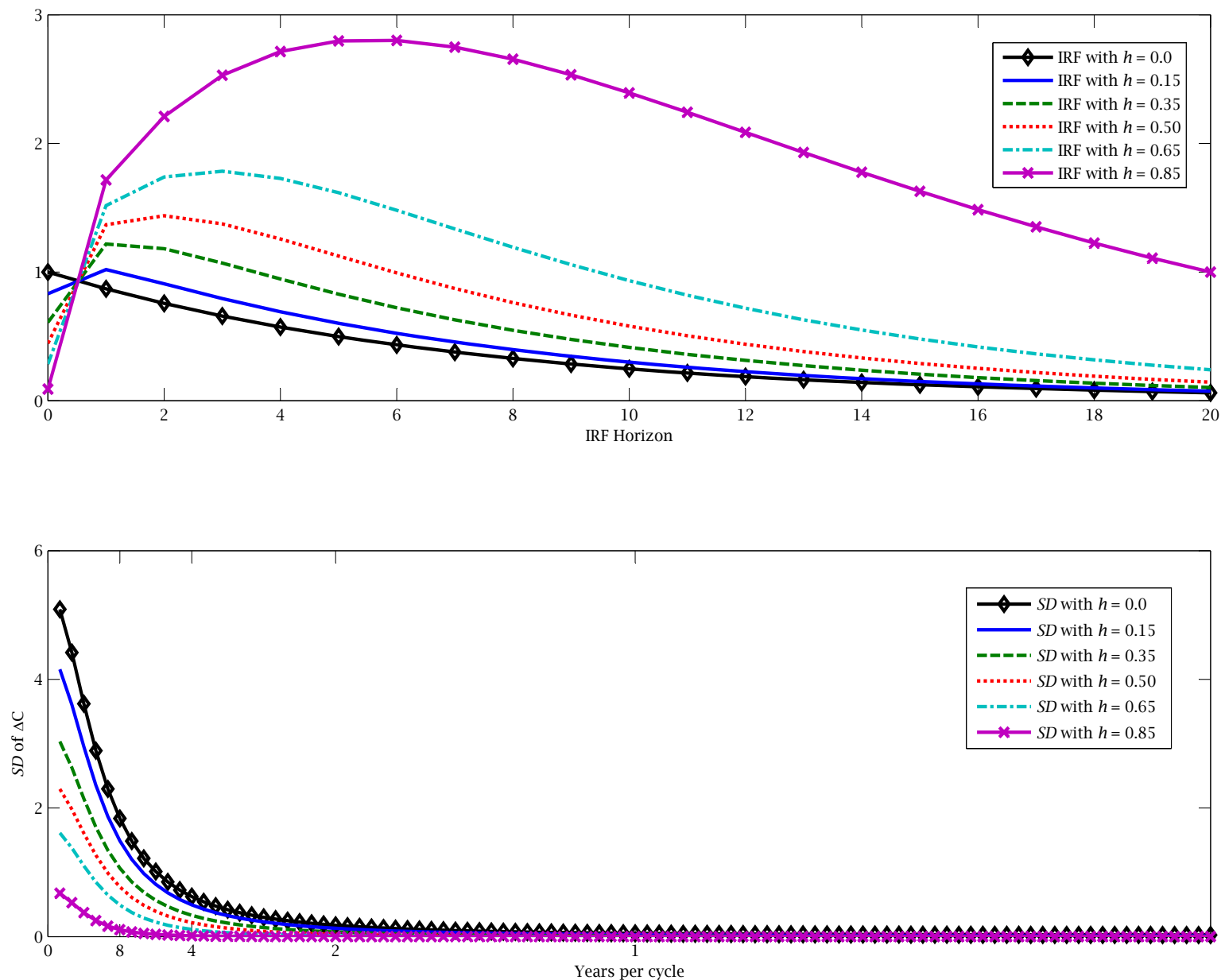
The calibration relies on existing DSGE model literature; see the text for details. For a non-informative prior, the right-most column contains the lower and upper end points of the uniform distribution. When the prior is based on the beta distribution, its two parameters are $a = \bar{\Gamma}_{i,n} [(1 - \bar{\Gamma}_{i,n})\bar{\Gamma}_{i,n}/STD(\Gamma_{i,n})^2 - 1]$ and $b = a(1 - \bar{\Gamma}_{i,n})/\bar{\Gamma}_{i,n}$, where $\bar{\Gamma}_{i,n}$ is the degenerate prior of the i th element of the parameter vector of model $n = 1, \dots, 4$, and its standard deviation is $STD(\Gamma_{i,n})$.

TABLE 2: CIC OF KOLMOGOROV-SMIRNOV STATISTICS

Model	ΔY w/r/t Trend Shock		ΔY w/r/t Transitory Shock		ΔC w/r/t Trend Shock		ΔC w/r/t Transitory Shock	
	$[0, \pi]$	$[8, 2]$	$[0, \pi]$	$[8, 2]$	$[0, \pi]$	$[8, 2]$	$[0, \pi]$	$[8, 2]$
Non-Habit NKDSGE								
Baseline								
MGR	0.02	0.03	0.00	0.01	0.00	0.00	0.00	0.00
TR	0.01	0.00	0.12	0.71	0.00	0.00	0.08	0.68
SPrice								
MGR	0.03	0.47	0.00	0.23	0.01	0.17	0.00	0.04
TR	0.40	0.57	0.00	0.76	0.01	0.16	0.00	0.49
SWage								
MGR	0.00	0.00	0.00	0.06	0.00	0.00	0.00	0.06
TR	0.00	0.00	0.21	0.37	0.00	0.00	0.02	0.81
Habit NKDSGE								
Baseline								
MGR	0.00	0.04	0.16	0.18	0.02	0.16	0.13	0.18
TR	0.00	0.03	0.64	0.52	0.03	0.14	0.53	0.85
SPrice								
MGR	0.14	0.64	0.11	0.59	0.09	0.44	0.29	0.49
TR	0.43	0.74	0.29	0.65	0.15	0.46	0.33	0.76
SWage								
MGR	0.00	0.01	0.18	0.23	0.01	0.10	0.13	0.23
TR	0.00	0.05	0.55	0.45	0.03	0.13	0.44	0.77

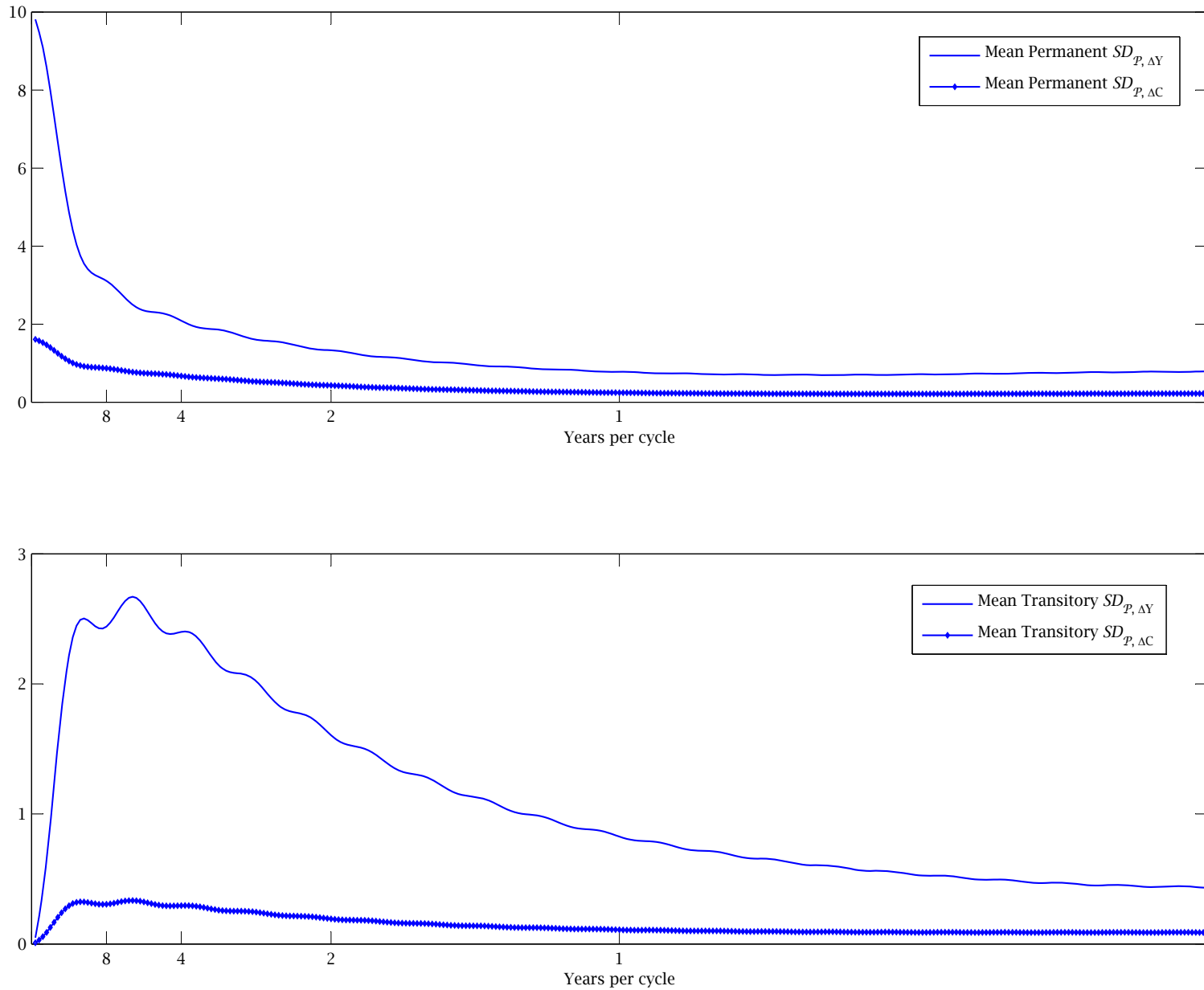
Baseline NKDSGE models include sticky prices and sticky wages. The acronyms SPrice and SWage represent NKDSGE models with only sticky prices or sticky nominal wages, respectively. The money growth and Taylor rules of equations (8) and (9) are denoted by MGR and the TR, respectively. The column headings $[0, \pi]$ and $[8, 2]$ indicate that the *CIC* quantifies the intersection of KS_P and KS_T distributions computed from permanent and transitory *SDs* of ΔY and ΔC with domains on the entire spectrum (*i.e.*, from frequency zero or the long run to the short run), and from eight to two years per cycle, respectively.

FIGURE 1: ΔC AND A REAL INTEREST RATE SHOCK: IMPULSE RESPONSE FUNCTIONS AND SDs



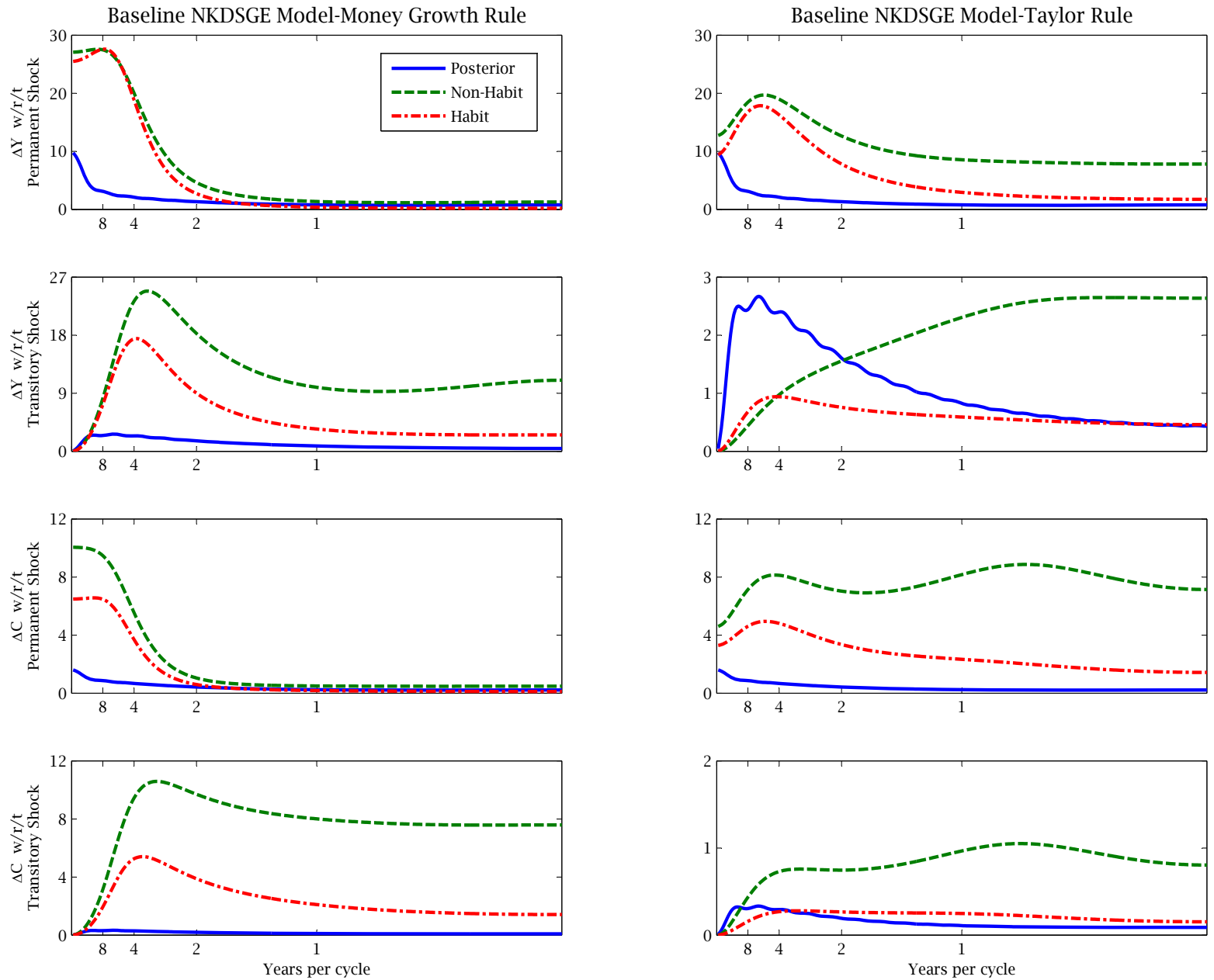
The top (bottom) window plots the impulse response functions (SDs) of ΔC generated from the solved linearized Euler (2) given a 1 percent shock to the forecast innovation of the AR(1) of the real rate, q_t .

FIGURE 2: MEAN PERMANENT AND TRANSITORY $SD_{\mathcal{P}}S$



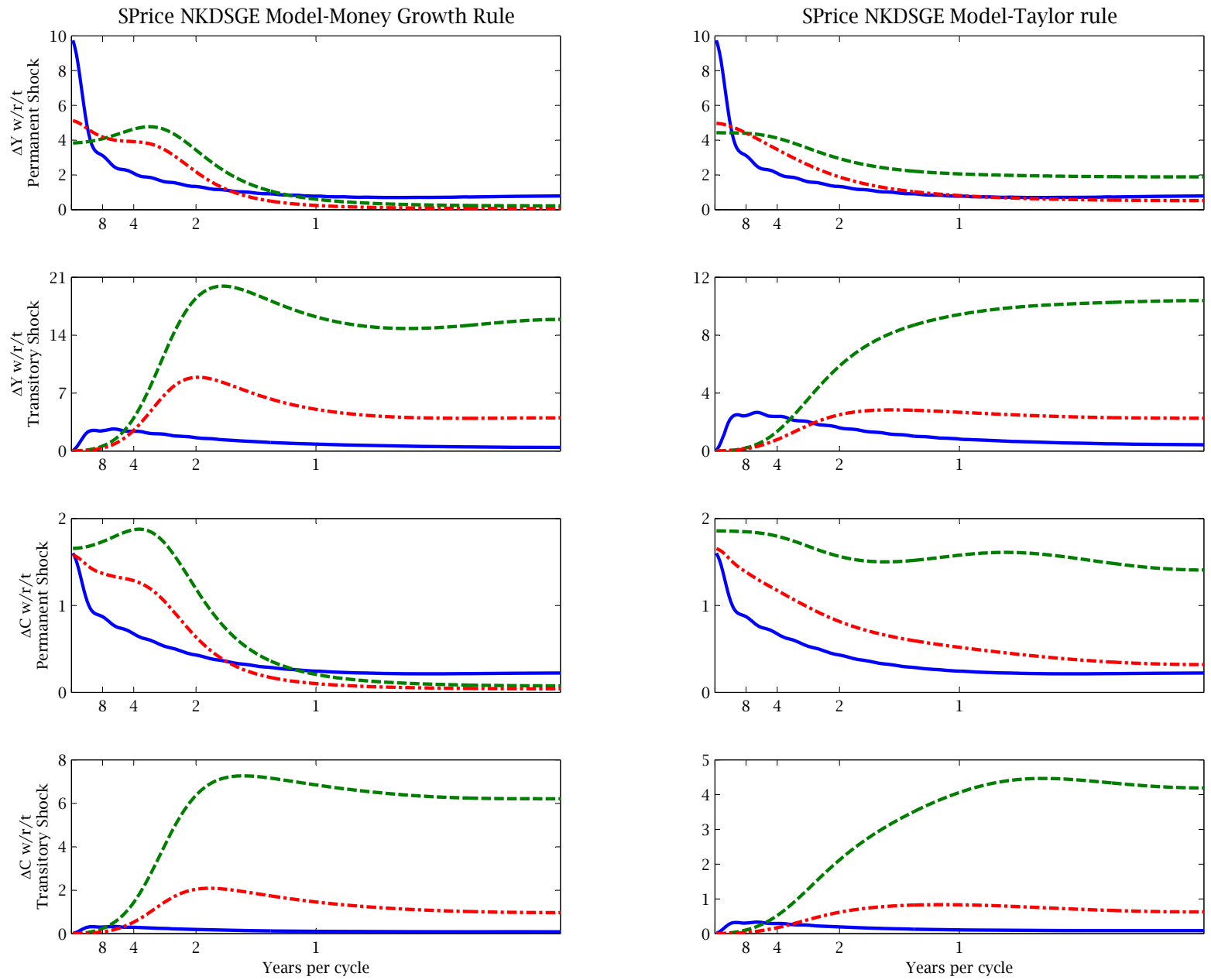
Mean permanent and transitory $SD_{\mathcal{P}, \Delta Y}$ and $SD_{\mathcal{P}, \Delta C}$ are averaged frequency by frequency across ensembles that consist of J of these SD s. The SD s are constructed using $SVMA(\infty)$ s that rely on LRMN, the BQ decomposition, and unrestricted VAR(2)s.

FIGURE 3: MEAN STRUCTURAL SD_{PS} AND SD_{TS} OF BASELINE NKDSGE MODELS



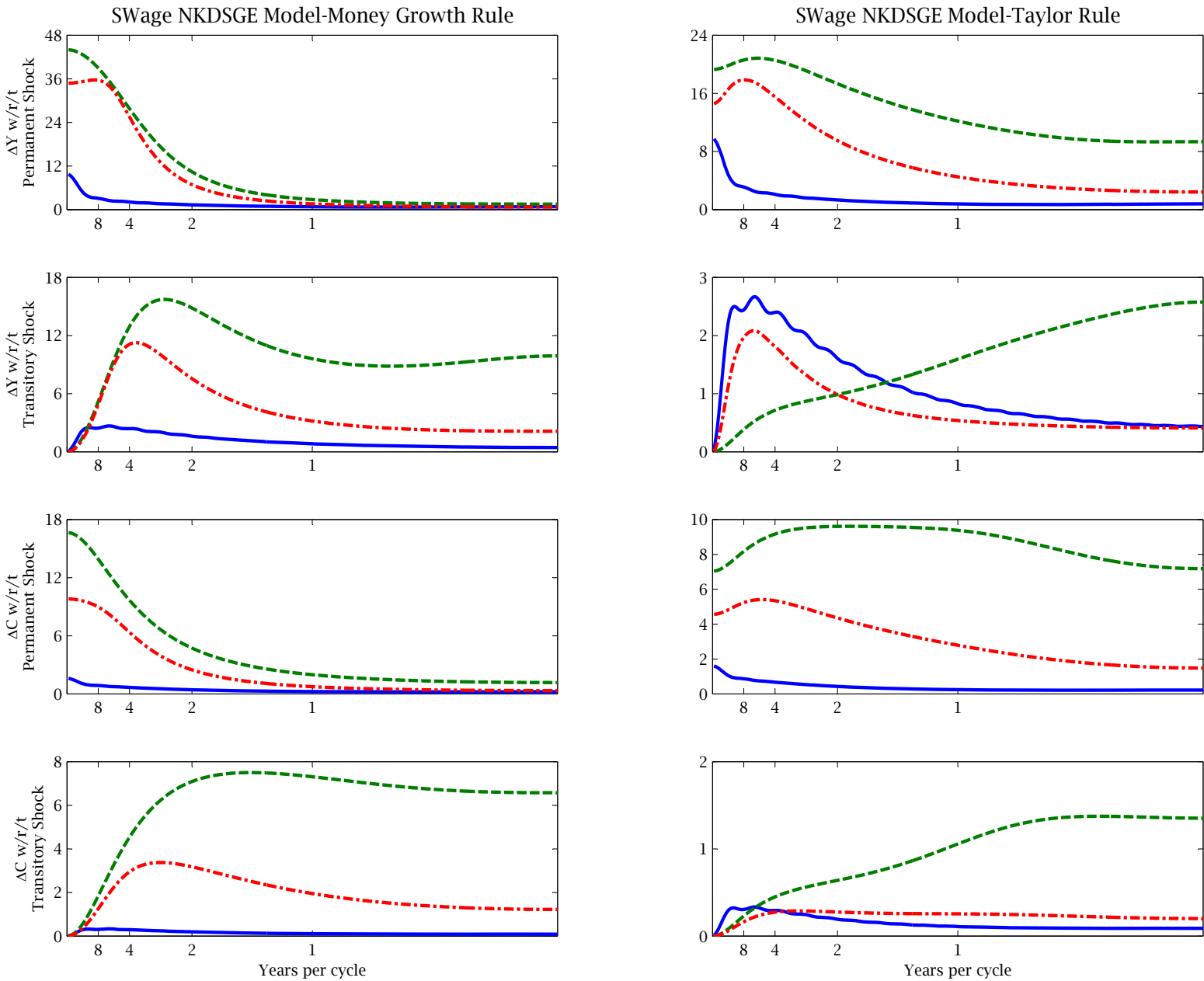
The habit and non-habit NKDSGE models generate ensembles of permanent and transitory $SD_{T,\Delta Y}$ s and $SD_{T,\Delta C}$ s that are averaged frequency by frequency to produce the mean SD_{TS} s. Otherwise, see the notes to figure 2.

FIGURE 4: MEAN STRUCTURAL SD_P S AND SD_T S OF SPRICE NKDSGE MODELS



See figure 3 for details.

FIGURE 5: MEAN STRUCTURAL SD_P S AND SD_T S OF SWAGE NKDSGE MODELS



See figure 3 for details.